# Robust Statistics and Generative Adversarial Networks

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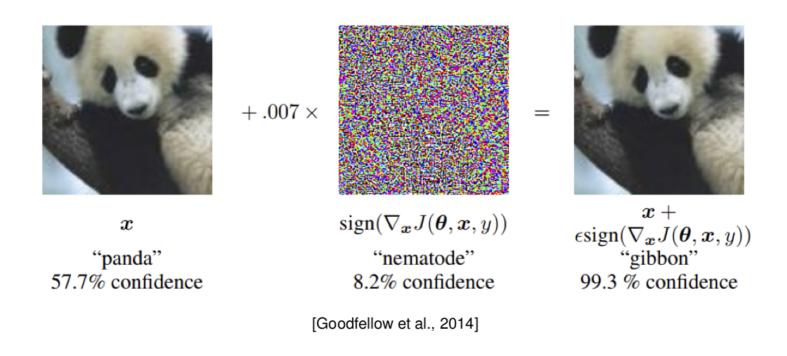




Chao Gao (Chicago) Jiyu Liu (Yale)

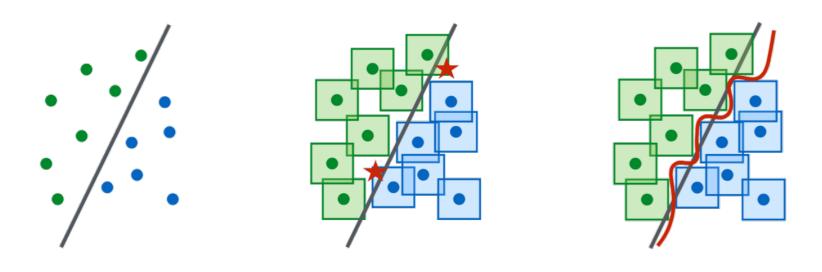
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## Deep Learning is Notoriously Not Robust!



- Imperceivable adversarial examples are ubiquitous to fail neural networks
- How can one achieve robustness?

# Robust Optimization



• Traditional training:

$$\min_{\theta} J_n(\theta, \mathbf{z} = (x_i, y_i)_{i=1}^n)$$

- e.g. square or cross-entropy loss as negative log-likelihood of logit models
- Robust optimization (Madry et al. ICLR'2018):

$$\min_{\theta} \max_{\|\epsilon_i\| \leq \delta} J_n(\theta, \mathbf{z} = (x_i + \epsilon_i, y_i)_{i=1}^n)$$

robust to any distributions, yet computationally hard

# Distributionally Robust Optimization (DRO)

• Distributional Robust Optimization:

$$\min_{\theta} \max_{\epsilon} \mathbb{E}_{\mathbf{z} \sim P_{\epsilon} \in \mathcal{D}}[J_n(\theta, \mathbf{z})]$$

ullet  $\mathcal D$  is a set of ambiguous distributions, e.g. Wasserstein ambiguity set

$$\mathcal{D} = \{P_{\epsilon} : W_2(P_{\epsilon}, \text{uniform distribution}) \leq \epsilon\}$$

where DRO may be reduced to regularized maximum likelihood estimates (Shafieezadeh-Abadeh, Esfahani, Kuhn, NIPS'2015) that are convex optimizations and tractable

# Wasserstein DRO and Sqrt-Lasso

Theorem (B., Kang, Murthy (2016)) Suppose that

$$c\left(\left(x,y\right),\left(x',y'\right)\right) = \begin{cases} \left\|x-x'\right\|_{q}^{2} & \text{if } y=y'\\ \infty & \text{if } y\neq y' \end{cases}.$$

Then, if 1/p + 1/q = 1

$$\max_{P:D_{c}(P,P_{n})\leq\delta}E_{P}^{1/2}\left(\left(Y-\beta^{T}X\right)^{2}\right)=E_{P_{n}}^{1/2}\left[\left(Y-\beta^{T}X\right)^{2}\right]+\sqrt{\delta}\left\|\beta\right\|_{p}.$$

Remark 1: This is sqrt-Lasso (Belloni et al. (2011)).

**Remark 2:** Uses RoPA duality theorem & "judicious choice of  $c(\cdot)$ "

#### Certified Robustness of Lasso

Take  $q = \infty$  and p = 1, with

$$c\left((x,y),\left(x',y'\right)\right) = \begin{cases} \|x - x'\|_{\infty}^{2} & \text{if } y = y' \\ \infty & \text{if } y \neq y' \end{cases}$$

Then for

$$P_n' = \frac{1}{n} \sum_i \delta_{x_i'}$$

with  $||x_i - x_i'||_{\infty} \leq \delta$ ,

$$D_c(P'_n,P_n)=\int \pi((x,y),(x',y'))c\left((x,y),(x',y')\right)\leq \delta,$$

for small enough  $\delta$  and well-separated x's. Sqrt-Lasso

$$\min_{\beta} \left\{ E_{P_n}^{1/2} \left[ \left( Y - \beta^T X \right)^2 \right] + \sqrt{\delta} \|\beta\|_1 \right\}^2$$

$$= \min_{\beta} \max_{P: D_c(P, P_n) \le \delta} E_P \left( \left( Y - \beta^T X \right)^2 \right)$$

provides a certified robust estimate in terms of Madry's adversarial training, using a convex Wasserstein relaxation.

## TV-neighborhood

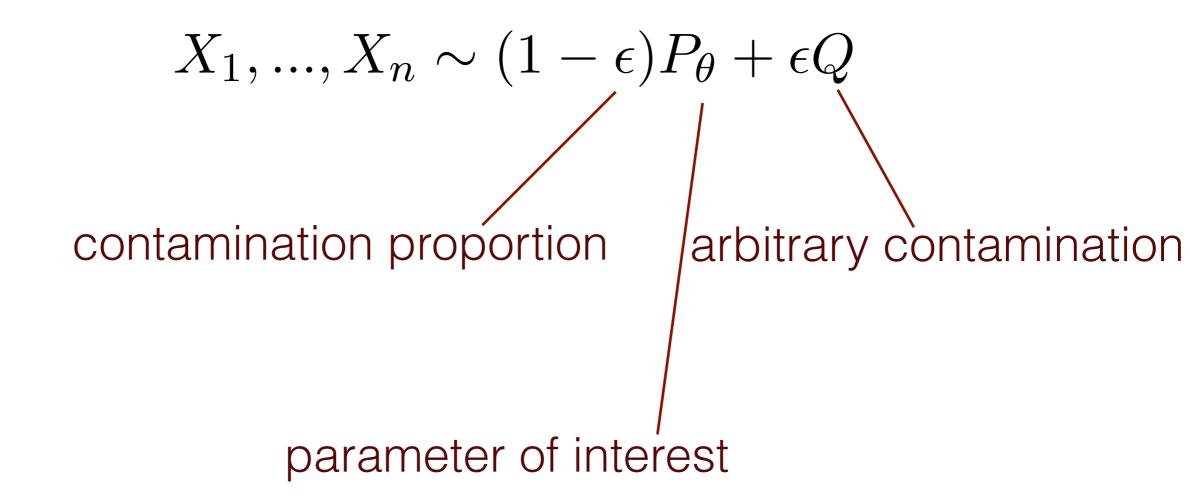
Now how about the TV-uncertainty set?

```
\mathcal{D} = \{P_{\epsilon} : TV(P_{\epsilon}, \text{uniform distribution}) \leq \epsilon\}?
```

$$X_1, ..., X_n \sim (1 - \epsilon)P_\theta + \epsilon Q$$

$$X_1,...,X_n \sim (1-\epsilon)P_{ heta} + \epsilon Q$$
 parameter of interest

$$X_1,...,X_n \sim (1-\epsilon)P_\theta + \epsilon Q$$
 contamination proportion parameter of interest



## An Example

$$X_1,...,X_n \sim (1-\epsilon)N(\theta,I_p) + \epsilon Q.$$

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how to estimate?

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$$\ell( heta,Q) = ext{negative log-likelihood} = \sum_{i=1}^{"} ( heta-X_i)^2 \ \sim (1-\epsilon)\mathbb{E}_{\mathcal{N}( heta)}( heta-X)^2 + \epsilon\mathbb{E}_Q( heta-X)^2$$

the sample mean

$$\hat{\theta}_{mean} = \frac{1}{n} \sum_{i=1}^{n} X_i = \arg\min_{\theta} \ell(\theta, Q)$$

$$\min_{\theta} \max_{Q} \ell(\theta, Q) \geq \max_{Q} \min_{\theta} \ell(\theta, Q) = \max_{Q} \ell(\hat{\theta}_{mean}, Q) = \infty$$

#### Medians

#### 1. Coordinatewise median

$$\hat{\theta} = (\hat{\theta}_j)$$
, where  $\hat{\theta}_j = \text{Median}(\{X_{ij}\}_{i=1}^n)$ ;

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#### 2. Tukey's median

$$\hat{\theta} = \arg\max_{\eta \in \mathbb{R}^p} \min_{||u||=1} \frac{1}{n} \sum_{i=1}^n \mathbb{I}\{u^T X_i > u^T \eta\}.$$

# Comparisons

	Coordinatewise Median	Tukey's Median
breakdown point	1/2	1/3
statistical precision	$\frac{p}{n}$	$\frac{p}{n}$
(no contamination)		
statistical precision	$\frac{p}{n} + p\epsilon^2$	$\frac{p}{n} + \epsilon^2$ : minimax
(with contamination)	<b>''</b>	[Chen-Gao-Ren'15]
computational complexity	Polynomial	NP-hard
		[Amenta et al. '00]

Note: R-package for Tukey median can not deal with more than 10 dimensions!

[https://github.com/ChenMengjie/DepthDescent]

$$\left\{ \frac{1}{n} \sum_{i=1}^{n} \mathbb{I}\{u^{T} X_{i} > u^{T} \eta\} \wedge \frac{1}{n} \sum_{i=1}^{n} \mathbb{I}\{u^{T} X_{i} \leq u^{T} \eta\} \right\}$$

$$\min_{\|u\|=1} \left\{ \frac{1}{n} \sum_{i=1}^{n} \mathbb{I}\{u^{T} X_{i} > u^{T} \eta\} \wedge \frac{1}{n} \sum_{i=1}^{n} \mathbb{I}\{u^{T} X_{i} \leq u^{T} \eta\} \right\}$$

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$$= \arg\max_{\eta \in \mathbb{R}^p} \min_{\|u\|=1} \frac{1}{n} \sum_{i=1}^n \mathbb{I}\{u^T X_i > u^T \eta\}.$$

model

$$y|X \sim N(X^T \beta, \sigma^2)$$

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$$u^T X y | X \sim N(u^T X X^T \beta, \sigma^2 u^T X X^T u)$$

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$$\left\{ \frac{1}{n} \sum_{i=1}^{n} \mathbb{I}\{u^{T} X_{i}(y_{i} - X_{i}^{T} \eta) > 0\} \wedge \frac{1}{n} \sum_{i=1}^{n} \mathbb{I}\{u^{T} X_{i}(y_{i} - X_{i}^{T} \eta) \leq 0\} \right\}$$

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$$\min_{u \in \mathbb{R}^p} \left\{ \frac{1}{n} \sum_{i=1}^n \mathbb{I} \{ u^T X_i (y_i - X_i^T \eta) > 0 \} \wedge \frac{1}{n} \sum_{i=1}^n \mathbb{I} \{ u^T X_i (y_i - X_i^T \eta) \le 0 \} \right\}$$

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$$\hat{\beta} = \underset{\eta \in \mathbb{R}^p}{\operatorname{argmax}} \min_{u \in \mathbb{R}^p} \left\{ \frac{1}{n} \sum_{i=1}^n \mathbb{I} \{ u^T X_i (y_i - X_i^T \eta) > 0 \} \wedge \frac{1}{n} \sum_{i=1}^n \mathbb{I} \{ u^T X_i (y_i - X_i^T \eta) \le 0 \} \right\}$$

model

$$y|X \sim N(X^T\beta, \sigma^2)$$

embedding

$$Xy|X \sim N(XX^T\beta, \sigma^2 XX^T)$$

projection

$$u^T X y | X \sim N(u^T X X^T \beta, \sigma^2 u^T X X^T u)$$

$$\hat{\beta} = \underset{\eta \in \mathbb{R}^p}{\operatorname{argmax}} \min_{u \in \mathbb{R}^p} \left\{ \frac{1}{n} \sum_{i=1}^n \mathbb{I}\{u^T X_i (y_i - X_i^T \eta) > 0\} \wedge \frac{1}{n} \sum_{i=1}^n \mathbb{I}\{u^T X_i (y_i - X_i^T \eta) \le 0\} \right\}$$

[Rousseeuw & Hubert, 1999]

# Tukey's depth is not a special case of regression depth.

$$(X,Y) \in \mathbb{R}^p \times \mathbb{R}^m \sim \mathbb{P}$$

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#### population version:

$$\mathcal{D}_{\mathcal{U}}(B, \mathbb{P}) = \inf_{U \in \mathcal{U}} \mathbb{P} \left\{ \left\langle U^T X, Y - B^T X \right\rangle \ge 0 \right\}$$

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#### empirical version:

$$\mathcal{D}_{\mathcal{U}}(B, \{(X_i, Y_i)\}_{i=1}^n) = \inf_{U \in \mathcal{U}} \frac{1}{n} \sum_{i=1}^n \mathbb{I} \left\{ \left\langle U^T X_i, Y_i - B^T X_i \right\rangle \ge 0 \right\}$$

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$$p = 1, X = 1 \in \mathbb{R},$$
 
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$$m = 1,$$

$$\mathcal{D}_{\mathcal{U}}(\beta, \mathbb{P}) = \inf_{U \in \mathcal{U}} \mathbb{P} \left\{ u^T X (y - \beta^T X) \ge 0 \right\}$$

**Proposition.** For any  $\delta > 0$ ,

$$\sup_{B \in \mathbb{R}^{p \times m}} |\mathcal{D}(B, \mathbb{P}_n) - \mathcal{D}(B, \mathbb{P})| \le C\sqrt{\frac{pm}{n}} + \sqrt{\frac{\log(1/\delta)}{2n}},$$

with probability at least  $1-2\delta$ .

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with probability at least  $1-2\delta$ .

#### Proposition.

$$\sup_{B,Q} |\mathcal{D}(B, (1 - \epsilon P_{B^*}) + \epsilon Q) - \mathcal{D}(B, P_{B^*})| \le \epsilon$$

 $(X,Y) \sim P_B$ 

$$(X,Y) \sim P_B: X \sim N(0,\Sigma), \quad Y|X \sim N(B^T X, \sigma^2 I_m)$$

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 $(X_1, Y_1), ..., (X_n, Y_n) \sim (1 - \epsilon) P_B + \epsilon Q$ 

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$$(X_1, Y_1), ..., (X_n, Y_n) \sim (1 - \epsilon)P_B + \epsilon Q$$

#### Theorem [G17]. For some C > 0,

$$\operatorname{Tr}((\widehat{B}-B)^T\Sigma(\widehat{B}-B)) \leq C\sigma^2\left(\frac{pm}{n}\vee\epsilon^2\right),$$

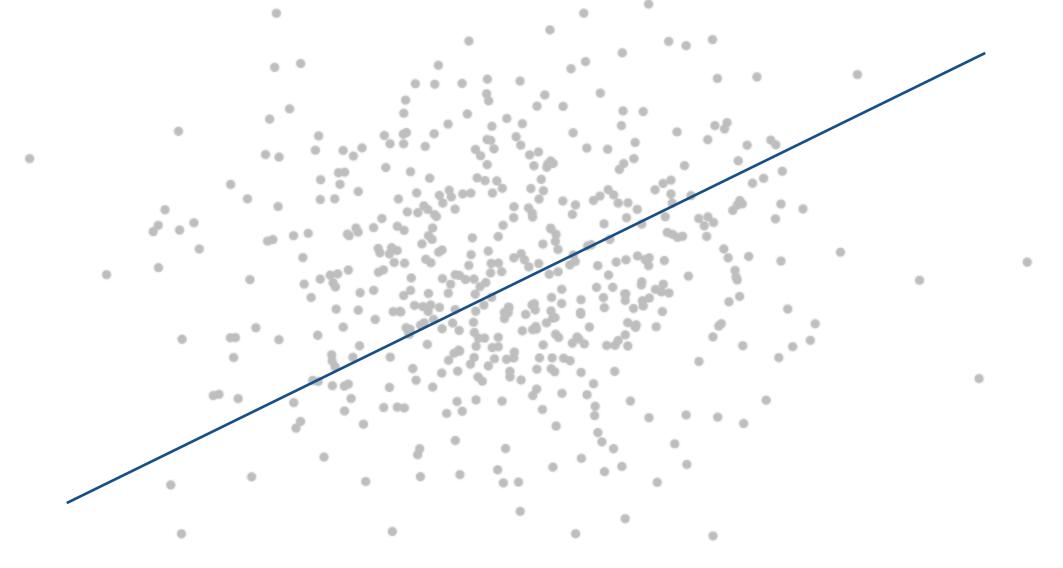
$$\|\widehat{B} - B\|_{\mathrm{F}}^2 \le C \frac{\sigma^2}{\kappa^2} \left( \frac{pm}{n} \vee \epsilon^2 \right),$$

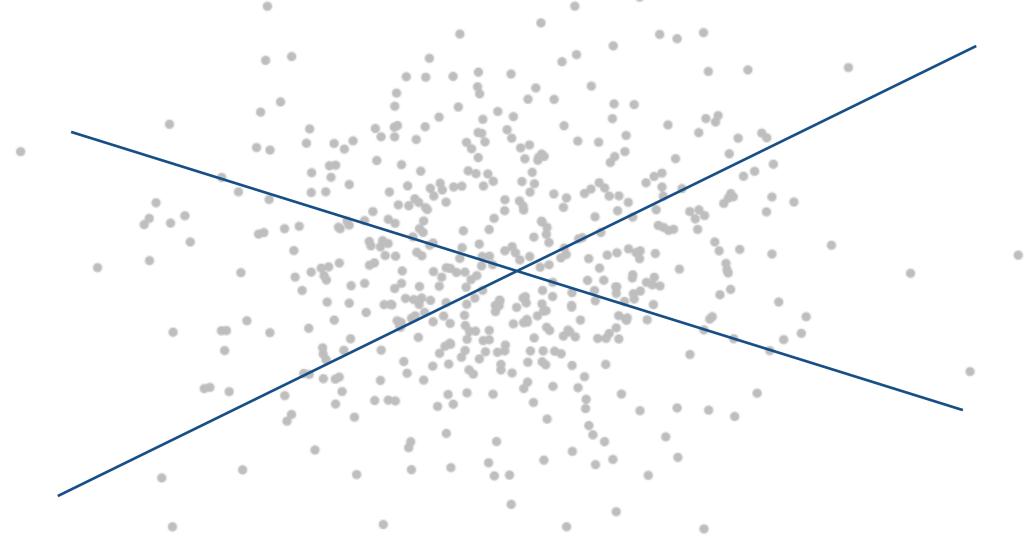
with high probability uniformly over B,Q.

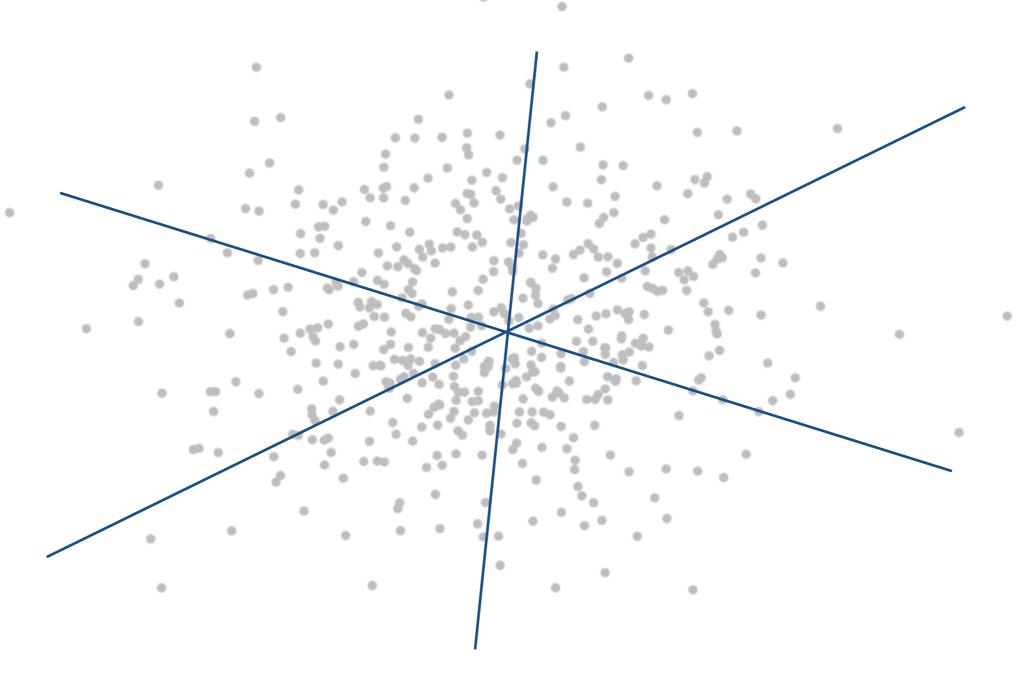
$$X_1,...,X_n \sim (1-\epsilon)N(0,\Sigma) + \epsilon Q.$$

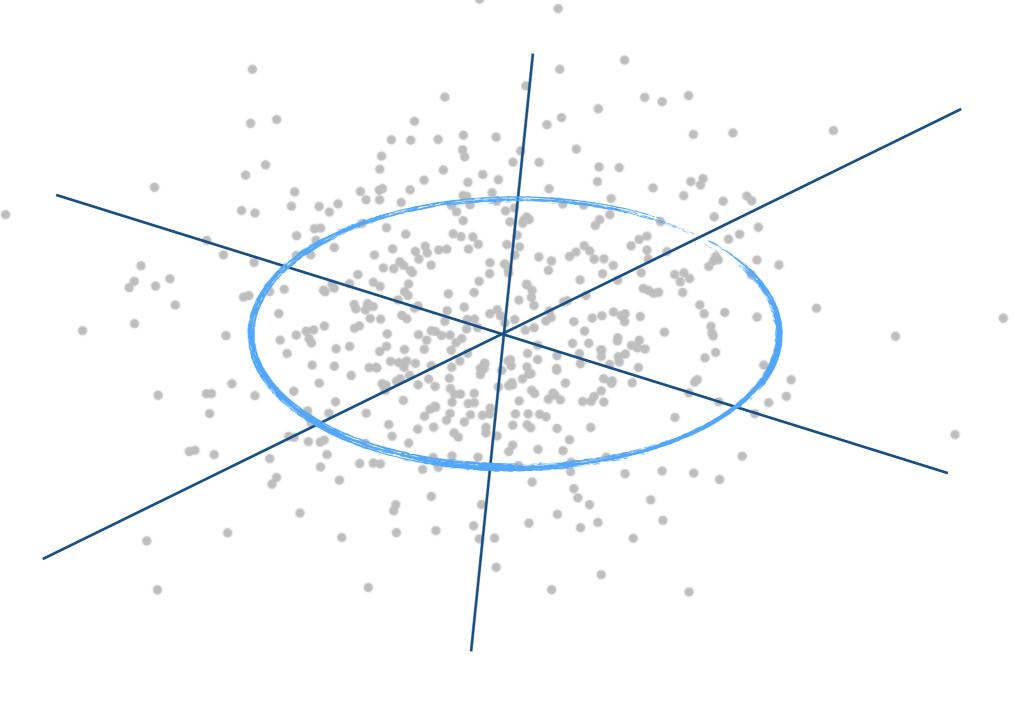
$$X_1, ..., X_n \sim (1 - \epsilon)N(0, \Sigma) + \epsilon Q.$$

how to estimate?









$$\mathcal{D}(\Gamma, \{X_i\}_{i=1}^n) = \min_{\|u\|=1} \min \left\{ \frac{1}{n} \sum_{i=1}^n \mathbb{I}\{|u^T X_i|^2 \ge u^T \Gamma u\}, \frac{1}{n} \sum_{i=1}^n \mathbb{I}\{|u^T X_i|^2 < u^T \Gamma u\} \right\}$$

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$$\hat{\Gamma} = \arg \max_{\Gamma \succeq 0} \mathcal{D}(\Gamma, \{X_i\}_{i=1}^n) \qquad \hat{\Sigma} = \hat{\Gamma}/\beta$$

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$$\hat{\Gamma} = \arg \max_{\Gamma \succeq 0} \mathcal{D}(\Gamma, \{X_i\}_{i=1}^n) \qquad \hat{\Sigma} = \hat{\Gamma}/\beta$$

Theorem [CGR15]. For some C > 0,

$$\|\hat{\Sigma} - \Sigma\|_{\text{op}}^2 \le C\left(\frac{p}{n} \vee \epsilon^2\right)$$

with high probability uniformly over  $\Sigma, Q$ .

## Summary

mean	$\ \cdot\ ^2$	$\frac{p}{n} \vee \epsilon^2$	
reduced rank regression	$\lVert \cdot \rVert_{ ext{F}}^2$	$\frac{\sigma^2}{\kappa^2} \frac{r(p+m)}{n} \vee \frac{\sigma^2}{\kappa^2} \epsilon^2$	
Gaussian graphical model	$\ \cdot\ _{\ell_1}^2$	$\frac{s^2 \log(ep/s)}{n} \vee s\epsilon^2$	
covariance matrix	$\ \cdot\ _{\mathrm{op}}^2$	$\frac{p}{n} \vee \epsilon^2$	
sparse PCA	$\lVert \cdot \rVert_{ ext{F}}^2$	$\frac{s\log(ep/s)}{n\lambda^2}\vee\frac{\epsilon^2}{\lambda^2}$	

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## Computation

### Computational Challenges

$$X_1,...,X_n \sim (1-\epsilon)N(\theta,I_p) + \epsilon Q.$$

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Lai, Rao, Vempala Diakonikolas, Kamath, Kane, Li, Moitra, Stewart Balakrishnan, Du, Singh

- Polynomial algorithms are proposed [Diakonikolas et al.'16, Lai et al. 16]
   of minimax optimal statistical precision
  - needs information on second or higher order of moments
  - ullet some priori knowledge about  $\epsilon$

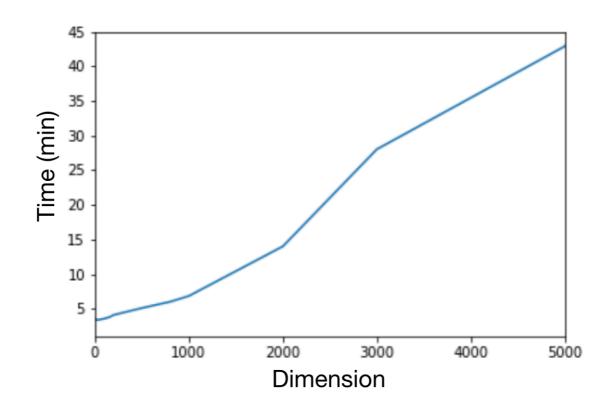
A well-defined objective function

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- Optimal for any elliptical distribution

### A practically good algorithm?

# Generative Adversarial Networks [Goodfellow et al. 2014]



Note: R-package for Tukey median can not deal with more than 10 dimensions [https://github.com/ChenMengjie/DepthDescent]

# Robust Learning of Cauchy Distributions

Table 4: Comparison of various methods of robust location estimation under Cauchy distributions. Samples are drawn from  $(1 - \epsilon)$ Cauchy $(0_p, I_p) + \epsilon Q$  with  $\epsilon = 0.2, p = 50$  and various choices of Q. Sample size: 50,000. Discriminator net structure: 50-50-25-1. Generator  $g_{\omega}(\xi)$  structure: 48-48-32-24-12-1 with absolute value activation function in the output layer.

Contamination $Q$	$JS$ - $GAN(G_1)$	$JS$ - $GAN(G_2)$	Dimension Halving	Iterative Filtering
Cauchy $(1.5 * 1_p, I_p)$	0.0664 (0.0065)	0.0743 (0.0103)	0.3529 (0.0543)	0.1244 (0.0114)
Cauchy $(5.0*1_p, I_p)$	0.0480 (0.0058)	0.0540 (0.0064)	0.4855 (0.0616)	0.1687 (0.0310)
Cauchy $(1.5*1_p, 5*I_p)$	0.0754 (0.0135)	0.0742 (0.0111)	0.3726 (0.0530)	0.1220 (0.0112)
$Normal(1.5*1_p, 5*I_p)$	0.0702 (0.0064)	0.0713 (0.0088)	0.3915 (0.0232)	0.1048 (0.0288))

- Dimension Halving: [Lai et al.'16] https://github.com/kal2000/AgnosticMeanAndCovarianceCode.
- Iterative Filtering: [Diakonikolas et al.'17]
  https://github.com/hoonose/robust-filter.

#### f-GAN

Given a strictly convex function f that satisfies f(1) = 0, the f-divergence between two probability distributions P and Q is defined by

$$D_f(P||Q) = \int f\left(\frac{p}{q}\right) dQ. \tag{8}$$

Let  $f^*$  be the convex conjugate of f. A variational lower bound of (8) is

$$D_f(P||Q) \ge \sup_{T \in \mathcal{T}} \left[ \mathbb{E}_P T(X) - \mathbb{E}_Q f^*(T(X)) \right]. \tag{9}$$

where equality holds whenever the class  $\mathcal{T}$  contains the function f'(p/q).

[Nowozin-Cseke-Tomioka'16] f-GAN minimizes the variational lower bound (9)

$$\widehat{P} = \underset{Q \in \mathcal{Q}}{\operatorname{arg \, min \, sup}} \left[ \frac{1}{n} \sum_{i=1}^{n} T(X_i) - \mathbb{E}_Q f^*(T(X)) \right]. \tag{10}$$

with i.i.d. observations  $X_1, ..., X_n \sim P$ .

# From f-GAN to Tukey's Median: f-learning

Consider the special case

$$\mathcal{T} = \left\{ f'\left(\frac{\widetilde{q}}{q}\right) : \widetilde{q} \in \widetilde{\mathcal{Q}} \right\}. \tag{11}$$

which is tight if  $P \in \widetilde{\mathcal{Q}}$ . The sample version leads to the following f-learning

$$\widehat{P} = \underset{Q \in \mathcal{Q}}{\operatorname{arg \, min \, sup}} \left[ \frac{1}{n} \sum_{i=1}^{n} f' \left( \frac{\widetilde{q}(X_i)}{q(X_i)} \right) - \mathbb{E}_Q f^* \left( f' \left( \frac{\widetilde{q}(X)}{q(X)} \right) \right) \right]. \tag{12}$$

- If  $f(x) = x \log x$ ,  $Q = \widetilde{Q}$ , (12)  $\Rightarrow$  Maximum Likelihood Estimate
- If f(x) = (x-1)+, then  $D_f(P||Q) = \frac{1}{2} \int |p-q|$  is the TV-distance,  $f^*(t) = t\mathbb{I}\{0 \le t \le 1\}$ ,  $f\text{-GAN} \Rightarrow \text{TV-GAN}$ 
  - $Q = \{N(\eta, I_p) : \eta \in \mathbb{R}^p\}$  and  $\widetilde{Q} = \{N(\widetilde{\eta}, I_p) : \|\widetilde{\eta} \eta\| \le r\}$ , (12)  $\stackrel{r \to 0}{\Rightarrow}$

Tukey's Median

## f-Learning

## f-Learning

f-divergence 
$$D_f(P||Q) = \int f\left(\frac{p}{q}\right) dQ$$

f-divergence 
$$D_f(P\|Q) = \int f\left(\frac{p}{q}\right) dQ$$

$$f(u) = \sup_{t} (tu - f^*(t))$$

f-divergence 
$$D_f(P||Q) = \int f\left(\frac{p}{q}\right)dQ$$

variational representation

$$= \sup_{T} \left[ \mathbb{E}_{X \sim P} T(X) - \mathbb{E}_{X \sim Q} f^*(T(X)) \right]$$

f-divergence 
$$D_f(P||Q) = \int f\left(\frac{p}{q}\right) dQ$$

#### variational representation

$$= \sup_{T} \left[ \mathbb{E}_{X \sim P} T(X) - \mathbb{E}_{X \sim Q} f^*(T(X)) \right]$$

$$T(x) = f'\left(\frac{p(x)}{q(x)}\right)$$

f-divergence 
$$D_f(P||Q) = \int f\left(\frac{p}{q}\right) dQ$$

variational representation

$$= \sup_{T} \left[ \mathbb{E}_{X \sim P} T(X) - \mathbb{E}_{X \sim Q} f^*(T(X)) \right]$$

$$= \sup_{\tilde{Q}} \left\{ \mathbb{E}_{X \sim P} f' \left( \frac{d\tilde{Q}(X)}{dQ(X)} \right) - \mathbb{E}_{X \sim Q} f^* \left( f' \left( \frac{d\tilde{Q}(X)}{dQ(X)} \right) \right) \right\}$$

$$\max_{T \in \mathcal{T}} \left\{ \frac{1}{n} \sum_{i=1}^{n} T(X_i) - \int f^*(T) dQ \right\}$$

$$\max_{\tilde{Q} \in \tilde{\mathcal{Q}}} \left\{ \frac{1}{n} \sum_{i=1}^{n} f' \left( \frac{\tilde{q}(X_i)}{q(X_i)} \right) - \int f^* \left( f' \left( \frac{\tilde{q}}{q} \right) \right) dQ \right\}$$

$$\min_{Q \in \mathcal{Q}} \max_{T \in \mathcal{T}} \left\{ \frac{1}{n} \sum_{i=1}^{n} T(X_i) - \int f^*(T) dQ \right\}$$

$$\min_{Q \in \mathcal{Q}} \max_{\tilde{Q} \in \tilde{\mathcal{Q}}} \left\{ \frac{1}{n} \sum_{i=1}^{n} f' \left( \frac{\tilde{q}(X_i)}{q(X_i)} \right) - \int f^* \left( f' \left( \frac{\tilde{q}}{q} \right) \right) dQ \right\}$$

**f-GAN** 
$$\min_{Q \in \mathcal{Q}} \max_{T \in \mathcal{T}} \left\{ \frac{1}{n} \sum_{i=1}^{n} T(X_i) - \int f^*(T) dQ \right\}$$

$$\textbf{f-Learning} \quad \min_{Q \in \mathcal{Q}} \max_{\tilde{Q} \in \tilde{\mathcal{Q}}} \left\{ \frac{1}{n} \sum_{i=1}^{n} f' \left( \frac{\tilde{q}(X_i)}{q(X_i)} \right) - \int f^* \left( f' \left( \frac{\tilde{q}}{q} \right) \right) dQ \right\}$$

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Jensen-Shannon	$f(x) = x \log x - (x+1)\log(x+1)$	GAN

[Goodfellow et al.]

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[Goodfellow et al., Baraud and Birge]

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Kullback-Leibler	$f(x) = x \log x$	MLE
Hellinger Squared	$f(x) = 2 - 2\sqrt{x}$	rho
<b>Total Variation</b>	$f(x) = (x-1)_+$	depth

[Goodfellow et al., Baraud and Birge]

$$\min_{Q \in \mathcal{Q}} \max_{\tilde{Q} \in \tilde{\mathcal{Q}}} \left\{ \frac{1}{n} \sum_{i=1}^{n} \mathbb{I} \left\{ \frac{\tilde{q}(X_i)}{q(X_i)} \ge 1 \right\} - Q \left( \frac{\tilde{q}}{q} \ge 1 \right) \right\}$$

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$$Q = \left\{ N(\theta, I_p) : \theta \in \mathbb{R}^p \right\} \qquad \tilde{Q} = \left\{ N(\tilde{\theta}, I_p) : \tilde{\theta} \in \mathcal{N}_r(\theta) \right\}$$

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Tukey depth  $\max_{\theta \in \mathbb{R}} \min_{\|u\|=1} \frac{1}{n} \sum_{i=1}^n \mathbb{I}\left\{u^T X_i \geq u^T \theta\right\}$ 

$$\min_{Q \in \mathcal{Q}} \max_{\tilde{Q} \in \tilde{\mathcal{Q}}} \left\{ \frac{1}{n} \sum_{i=1}^{n} \mathbb{I} \left\{ \frac{\tilde{q}(X_i)}{q(X_i)} \ge 1 \right\} - Q \left( \frac{\tilde{q}}{q} \ge 1 \right) \right\}$$

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$$\mathcal{Q} = \left\{ N(0, \Sigma) : \Sigma \in \mathbb{R}^{p \times p} \right\} \quad \tilde{\mathcal{Q}} = \left\{ N(0, \tilde{\Sigma}) : \tilde{\Sigma} = \Sigma + ruu^T, \|u\| = 1 \right\}$$

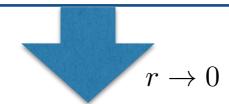
$$\min_{Q \in \mathcal{Q}} \max_{\tilde{Q} \in \tilde{\mathcal{Q}}} \left\{ \frac{1}{n} \sum_{i=1}^{n} \mathbb{I} \left\{ \frac{\tilde{q}(X_i)}{q(X_i)} \ge 1 \right\} - Q \left( \frac{\tilde{q}}{q} \ge 1 \right) \right\}$$

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# (related to) matrix depth

$$\max_{\Sigma} \min_{\|u\|=1} \left[ \left( \frac{1}{n} \sum_{i=1}^{n} \mathbb{I}\{|u^{T} X_{i}|^{2} \leq u^{T} \Sigma u\} - \mathbb{P}(\chi_{1}^{2} \leq 1) \right) \wedge \left( \frac{1}{n} \sum_{i=1}^{n} \mathbb{I}\{|u^{T} X_{i}|^{2} > u^{T} \Sigma u\} - \mathbb{P}(\chi_{1}^{2} > 1) \right) \right]$$

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practically good algorithms

#### theoretical foundation



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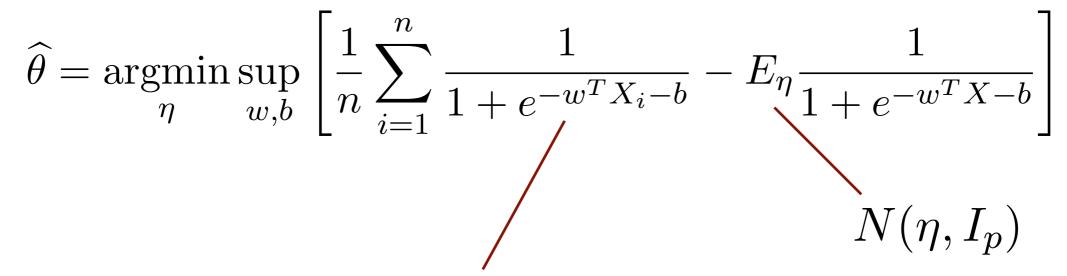


practically good algorithms

$$\widehat{\theta} = \underset{\eta}{\operatorname{argmin}} \sup_{w,b} \left[ \frac{1}{n} \sum_{i=1}^{n} \frac{1}{1 + e^{-w^{T} X_{i} - b}} - E_{\eta} \frac{1}{1 + e^{-w^{T} X - b}} \right]$$

$$\widehat{\theta} = \underset{\eta}{\operatorname{argmin}} \sup_{w,b} \left[ \frac{1}{n} \sum_{i=1}^{n} \frac{1}{1 + e^{-w^{T}X_{i} - b}} - E_{\eta} \frac{1}{1 + e^{-w^{T}X - b}} \right]$$

$$N(\eta, I_{p})$$



logistic regression classifier

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$$N(\eta, I_{p})$$

#### logistic regression classifier

#### Theorem [GLYZ18]. For some C > 0,

$$\|\widehat{\theta} - \theta\|^2 \le C\left(\frac{p}{n} \vee \epsilon^2\right)$$

with high probability uniformly over  $\theta \in \mathbb{R}^p, Q$ .

# TV-GAN rugged landscape!

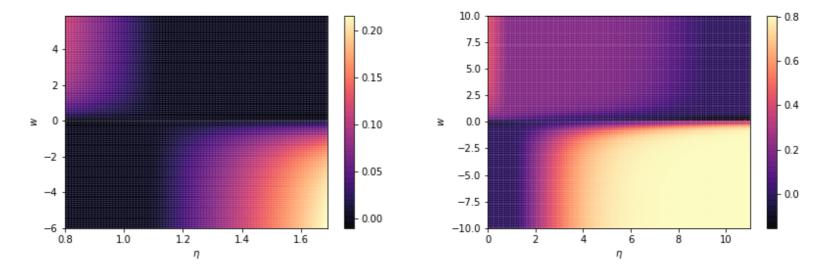


Figure: Heatmaps of the landscape of  $F(\eta,w)=\sup_b [E_P \operatorname{sigmoid}(wX+b)-E_{N(\eta,1)}\operatorname{sigmoid}(wX+b)]$ , where b is maximized out for visualization. Left: samples are drawn from  $P=(1-\epsilon)N(1,1)+\epsilon N(1.5,1)$  with  $\epsilon=0.2$ . Right: samples are drawn from  $P=(1-\epsilon)N(1,1)+\epsilon N(1.5,1)$  with  $\epsilon=0.2$ . Right: samples are drawn from  $P=(1-\epsilon)N(1,1)+\epsilon N(1.5,1)$  with  $\epsilon=0.2$ . Left: the landscape is good in the sense that no matter whether we start from the left-top area or the right-bottom area of the heatmap, gradient ascent on  $\eta$  does not consistently increase or decrease the value of  $\eta$ . This is because the signal becomes weak when it is close to the saddle point around  $\eta=1$ . Right: it is clear that  $\tilde{F}(w)=F(\eta,w)$  has two local maxima for a given  $\eta$ , achieved at  $w=+\infty$  and  $w=-\infty$ . In fact, the global maximum for  $\tilde{F}(w)$  has a phase transition from  $w=+\infty$  to  $w=-\infty$  as  $\eta$  grows. For example, the maximum is achieved at  $w=+\infty$  when  $\eta=1$  (blue solid) and is achieved at  $w=-\infty$  when  $\eta=5$  (red solid). Unfortunately, even if we initialize with  $\eta_0=1$  and  $w_0>0$ , gradient ascents on  $\eta$  will only increase the value of  $\eta$  (green dash), and thus as long as the discriminator cannot reach the global maximizer, w will be stuck in the positive half space  $\{w:w>0\}$  and further increase the value of  $\eta$ .

## The Original JS-GAN

[Goodfellow et al. 2014] For  $f(x) = x \log x - (x+1) \log \frac{x+1}{2}$ ,

$$\widehat{\theta} = \arg\min_{\eta \in \mathbb{R}^p} \max_{D \in \mathcal{D}} \left[ \frac{1}{n} \sum_{i=1}^n \log D(X_i) + \mathbb{E}_{\mathcal{N}(\eta, I_p)} \log(1 - D(X)) \right] + \log 4. \quad (15)$$

What are  $\mathcal{D}$ , the class of discriminators?

• Single layer (no hidden layer):

$$\mathcal{D} = \left\{ D(x) = \operatorname{sigmoid}(w^T x + b) : w \in \mathbb{R}^p, b \in \mathbb{R} \right\}$$

• One-hidden or Multiple layer:

$$\mathcal{D} = \left\{ D(x) = \operatorname{sigmoid}(w^T g(X)) \right\}$$

$$\widehat{\theta} = \underset{\eta \in \mathbb{R}^p}{\operatorname{argmin}} \max_{T \in \mathcal{T}} \left[ \frac{1}{n} \sum_{i=1}^n \log T(X_i) + E_{\eta} \log(1 - T(X)) \right] + \log 4$$

$$\widehat{\theta} = \underset{\eta \in \mathbb{R}^p}{\operatorname{argmin}} \max_{T \in \mathcal{T}} \left[ \frac{1}{n} \sum_{i=1}^n \log T(X_i) + E_{\eta} \log(1 - T(X)) \right] + \log 4$$

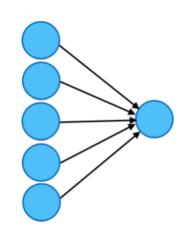
## numerical experiment

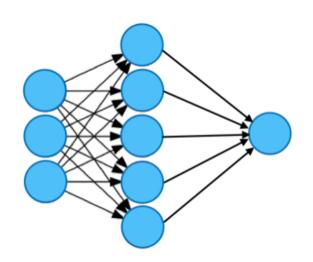
$$X_1, ..., X_n \sim (1 - \epsilon)N(\theta, I_p) + \epsilon N(\widetilde{\theta}, I_p)$$

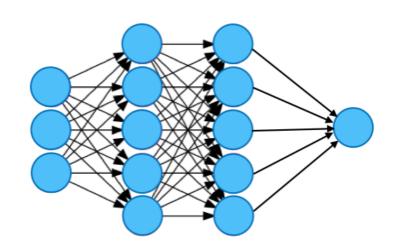
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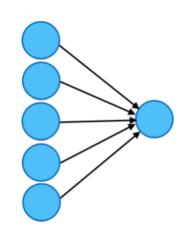


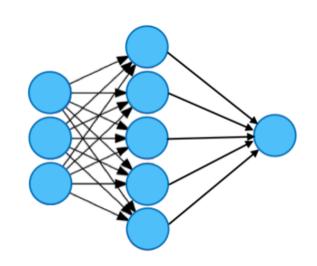


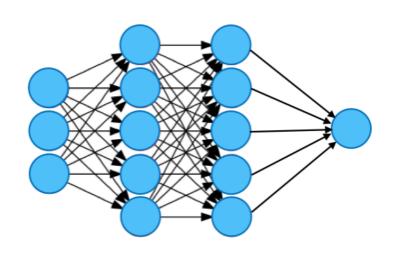
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## numerical experiment

$$X_1, ..., X_n \sim (1 - \epsilon)N(\theta, I_p) + \epsilon N(\widetilde{\theta}, I_p)$$







$$\widehat{\theta} \approx (1 - \epsilon)\theta + \epsilon \widetilde{\theta}$$



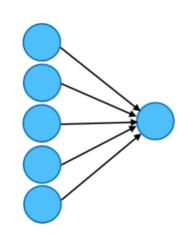




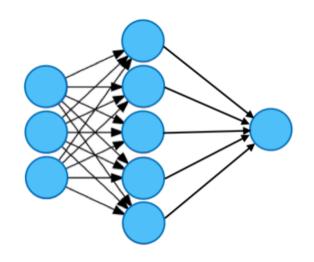
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#### numerical experiment

$$X_1, ..., X_n \sim (1 - \epsilon)N(\theta, I_p) + \epsilon N(\widetilde{\theta}, I_p)$$

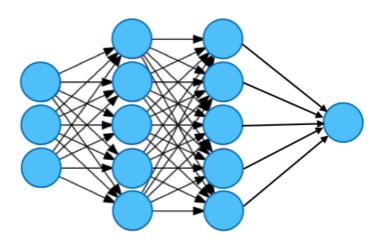


$$\widehat{\theta} \approx (1 - \epsilon)\theta + \epsilon \widetilde{\theta}$$



$$\widehat{\theta} \approx \theta$$





$$\widehat{\theta} \approx \theta$$







A classifier with hidden layers leads to robustness. Why?

#### A classifier with hidden layers leads to robustness. Why?

$$\mathsf{JS}_g(\mathbb{P},\mathbb{Q}) = \max_{w \in \mathbb{R}^d} \left[ \mathbb{P} \log \frac{1}{1 + e^{-w^T g(X)}} + \mathbb{Q} \log \frac{1}{1 + e^{w^T g(X)}} \right] + \log 4.$$

#### A classifier with hidden layers leads to robustness. Why?

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#### Proposition.

$$\mathsf{JS}_g(\mathbb{P},\mathbb{Q}) = 0 \iff \mathbb{P}g(X) = \mathbb{Q}g(X)$$

$$\widehat{\theta} = \underset{\eta \in \mathbb{R}^p}{\operatorname{argmin}} \max_{T \in \mathcal{T}} \left[ \frac{1}{n} \sum_{i=1}^n \log T(X_i) + E_{\eta} \log(1 - T(X)) \right] + \log 4$$

**Theorem [GLYZ18].** For a neural network class  $\mathcal{T}$  with at least one hidden layer and appropriate regularization, we have

$$\|\widehat{\theta} - \theta\|^2 \lesssim \begin{cases} \frac{p}{n} + \epsilon^2 & \text{(indicator/sigmoid/ramp)} \\ \frac{p \log p}{n} + \epsilon^2 & \text{(ReLUs+sigmoid features)} \end{cases}$$

with high probability uniformly over  $\theta \in \mathbb{R}^p, Q$ .

# JS-GAN: Adaptation to Unknown Covariance

unknown covariance?

$$X_1, ..., X_n \sim (1 - \epsilon)N(\theta, \Sigma) + \epsilon Q$$

# JS-GAN: Adaptation to Unknown Covariance

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$$(\widehat{\theta}, \widehat{\Sigma}) = \underset{\eta, \Gamma}{\operatorname{argmin}} \max_{T \in \mathcal{T}} \left[ \frac{1}{n} \sum_{i=1}^{n} \log T(X_i) + \mathbb{E}_{X \sim N(\eta, \Gamma)} \log(1 - T(X)) \right]$$

# JS-GAN: Adaptation to Unknown Covariance

### unknown covariance?

$$X_1, ..., X_n \sim (1 - \epsilon)N(\theta, \Sigma) + \epsilon Q$$

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no need to change the discriminator class

### Generalization

#### **Strong Contamination model:**

 $X_1, ..., X_n \stackrel{iid}{\sim} P$  for some P satisfying  $\mathsf{TV}(P, E(\theta, \Sigma, H)) \leq \epsilon$ 

$$(\widehat{\theta}, \widehat{\Sigma}, \widehat{H}) = \underset{\eta \in \mathbb{R}^p, \Gamma \in \mathcal{E}_p(M), H \in \mathcal{H}(M')}{\operatorname{argmin}} \max_{T \in \mathcal{T}} \left[ \frac{1}{n} \sum_{i=1}^n S(T(X_i), 1) + \mathbb{E}_{X \sim E(\eta, \Gamma, G)} S(T(X), 0) \right]$$

A scoring rule S is regular if both  $S(\cdot,0)$  and  $S(\cdot,1)$  are real-valued, except possibly that  $S(0,1) = -\infty$  or  $S(1,0) = -\infty$ . The celebrated Savage representation [50] asserts that a regular scoring rule S is proper if and only if there is a convex function  $G(\cdot)$ , such that

$$\begin{cases} S(t,1) = G(t) + (1-t)G'(t), \\ S(t,0) = G(t) - tG'(t). \end{cases}$$
(10)

Here, G'(t) is a subgradient of G at the point t. Moreover, the statement also holds for strictly proper scoring rules when convex is replaced by strictly convex.

### Consistency

**Theorem [GYZ19].** For a neural network class  $\mathcal{T}$  with at least one hidden layer and appropriate regularization, we have

$$\|\widehat{\theta} - \theta\|^2 \le C\left(\frac{p}{n} \vee \epsilon^2\right),$$
  
$$\|\widehat{\Sigma} - \Sigma\|_{\text{op}}^2 \le C\left(\frac{p}{n} \vee \epsilon^2\right),$$

## Example 1: Log Score and JS-GAN

1. Log Score. The log score is perhaps the most commonly used rule because of its various intriguing properties [31]. The scoring rule with  $S(t,1) = \log t$  and  $S(t,0) = \log(1-t)$  is regular and strictly proper. Its Savage representation is given by the convex function  $G(t) = t \log t + (1-t) \log(1-t)$ , which is interpreted as the negative Shannon entropy of Bernoulli(t). The corresponding divergence function  $D_{\mathcal{T}}(P,Q)$ , according to Proposition 3.1, is a variational lower bound of the Jensen-Shannon divergence

$$\mathsf{JS}(P,Q) = \frac{1}{2} \int \log \left( \frac{dP}{dP + dQ} \right) dP + \frac{1}{2} \int \log \left( \frac{dQ}{dP + dQ} \right) dQ + \log 2.$$

Its sample version (13) is the original GAN proposed by [25] that is widely used in learning distributions of images.

# Example 2: Zero-One Score and TV-GAN

2. Zero-One Score. The zero-one score  $S(t,1) = 2\mathbb{I}\{t \geq 1/2\}$  and  $S(t,0) = 2\mathbb{I}\{t < 1/2\}$  is also known as the misclassification loss. This is a regular proper scoring rule but not strictly proper. The induced divergence function  $D_{\mathcal{T}}(P,Q)$  is a variational lower bound of the total variation distance

$$\mathsf{TV}(P,Q) = P\left(\frac{dP}{dQ} \ge 1\right) - Q\left(\frac{dP}{dQ} \ge 1\right) = \frac{1}{2}\int |dP - dQ|.$$

The sample version (13) is recognized as the TV-GAN that is extensively studied by [21] in the context of robust estimation.

# Example 3: Quadratic Score and LS-GAN

3. Quadratic Score. Also known as the Brier score [6], the definition is given by  $S(t,1) = -(1-t)^2$  and  $S(t,0) = -t^2$ . The corresponding convex function in the Savage representation is given by G(t) = -t(1-t). By Proposition 2.1, the divergence function (3) induced by this regular strictly proper scoring rule is a variational lower bound of the following divergence function,

$$\Delta(P,Q) = \frac{1}{8} \int \frac{(dP - dQ)^2}{dP + dQ},$$

known as the triangular discrimination. The sample version (5) belongs to the family of least-squares GANs proposed by [39].

### Example 4: Boosting Score

4. Boosting Score. The boosting score was introduced by [7] with  $S(t,1) = -\left(\frac{1-t}{t}\right)^{1/2}$  and  $S(t,0) = -\left(\frac{t}{1-t}\right)^{1/2}$  and has an connection to the AdaBoost algorithm. The corresponding convex function in the Savage representation is given by  $G(t) = -2\sqrt{t(1-t)}$ . The induced divergence function  $D_{\mathcal{T}}(P,Q)$  is thus a variational lower bound of the squared Hellinger distance

$$H^2(P,Q) = \frac{1}{2} \int \left(\sqrt{dP} - \sqrt{dQ}\right)^2.$$

# Example 5: Beta Score and new GANs

5. Beta Score. A general Beta family of proper scoring rules was introduced by [7] with  $S(t,1) = -\int_t^1 c^{\alpha-1} (1-c)^{\beta} dc$  and  $S(t,0) = -\int_0^t c^{\alpha} (1-c)^{\beta-1} dc$  for any  $\alpha, \beta > -1$ . The log score, the quadratic score and the boosting score are special cases of the Beta score with  $\alpha = \beta = 0$ ,  $\alpha = \beta = 1$ ,  $\alpha = \beta = -1/2$ . The zero-one score is a limiting case of the Beta score by letting  $\alpha = \beta \to \infty$ . Moreover, it also leads to asymmetric scoring rules with  $\alpha \neq \beta$ .

# Robust Learning of Gaussian Distributions

Q	n	р	$\epsilon$	TV-GAN	JS-GAN	Dimension Halving	Iterative Filtering
$N(0.5*1_p,I_p)$	50,000	100	.2	0.0953 (0.0064)	0.1144 (0.0154)	0.3247 (0.0058)	0.1472 (0.0071)
$N(0.5*1_p,I_p)$	5,000	100	.2	0.1941 (0.0173)	0.2182 (0.0527)	0.3568 (0.0197)	0.2285 (0.0103)
$N(0.5*1_p,I_p)$	50,000	200	.2	0.1108 (0.0093)	0.1573 (0.0815)	0.3251 (0.0078)	0.1525 (0.0045)
$N(0.5*1_p,I_p)$	50,000	100	.05	0.0913 (0.0527)	0.1390 (0.0050)	0.0814 (0.0056)	0.0530 (0.0052)
$N(5*1_p,I_p)$	50,000	100	.2	2.7721 (0.1285)	0.0534 (0.0041)	0.3229 (0.0087)	0.1471 (0.0059)
$N(0.5*1_p,\Sigma)$	50,000	100	.2	0.1189 (0.0195)	0.1148 (0.0234)	0.3241 (0.0088)	0.1426 (0.0113)
$Cauchy(0.5*1_p)$	50,000	100	.2	0.0738 (0.0053)	0.0525 (0.0029)	0.1045 (0.0071)	0.0633 (0.0042)

Table: Comparison of various robust mean estimation methods. The smallest error of each case is highlighted in bold.

- Dimension Halving: [Lai et al.'16]

  https://github.com/kal2000/AgnosticMeanAndCovarianceCode.
- Iterative Filtering: [Diakonikolas et al.'17]
  https://github.com/hoonose/robust-filter.

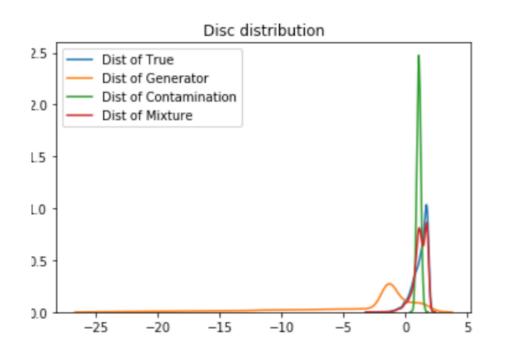
# Robust Learning of Cauchy Distributions

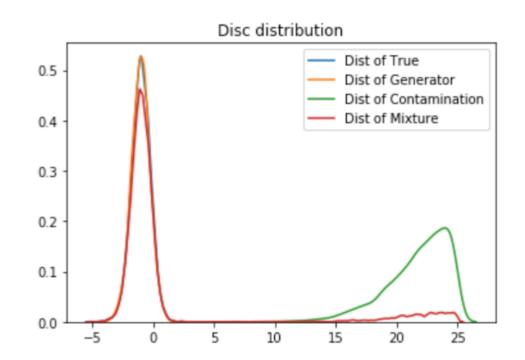
Table 4: Comparison of various methods of robust location estimation under Cauchy distributions. Samples are drawn from  $(1 - \epsilon)$ Cauchy $(0_p, I_p) + \epsilon Q$  with  $\epsilon = 0.2, p = 50$  and various choices of Q. Sample size: 50,000. Discriminator net structure: 50-50-25-1. Generator  $g_{\omega}(\xi)$  structure: 48-48-32-24-12-1 with absolute value activation function in the output layer.

Contamination $Q$	$JS$ - $GAN(G_1)$	$JS$ - $GAN(G_2)$	Dimension Halving	Iterative Filtering
Cauchy $(1.5*1_p, I_p)$	0.0664 (0.0065)	0.0743 (0.0103)	0.3529 (0.0543)	0.1244 (0.0114)
$Cauchy(5.0*1_p, I_p)$	0.0480 (0.0058)	0.0540 (0.0064)	0.4855 (0.0616)	0.1687 (0.0310)
Cauchy $(1.5*1_p, 5*I_p)$	0.0754 (0.0135)	0.0742 (0.0111)	0.3726 (0.0530)	0.1220 (0.0112)
Normal $(1.5*1_p, 5*I_p)$	0.0702 (0.0064)	0.0713 (0.0088)	0.3915 (0.0232)	0.1048 (0.0288))

- Dimension Halving: [Lai et al.'16] https://github.com/kal2000/AgnosticMeanAndCovarianceCode.
- Iterative Filtering: [Diakonikolas et al.'17] https://github.com/hoonose/robust-filter.

# Discriminator identifies outliers

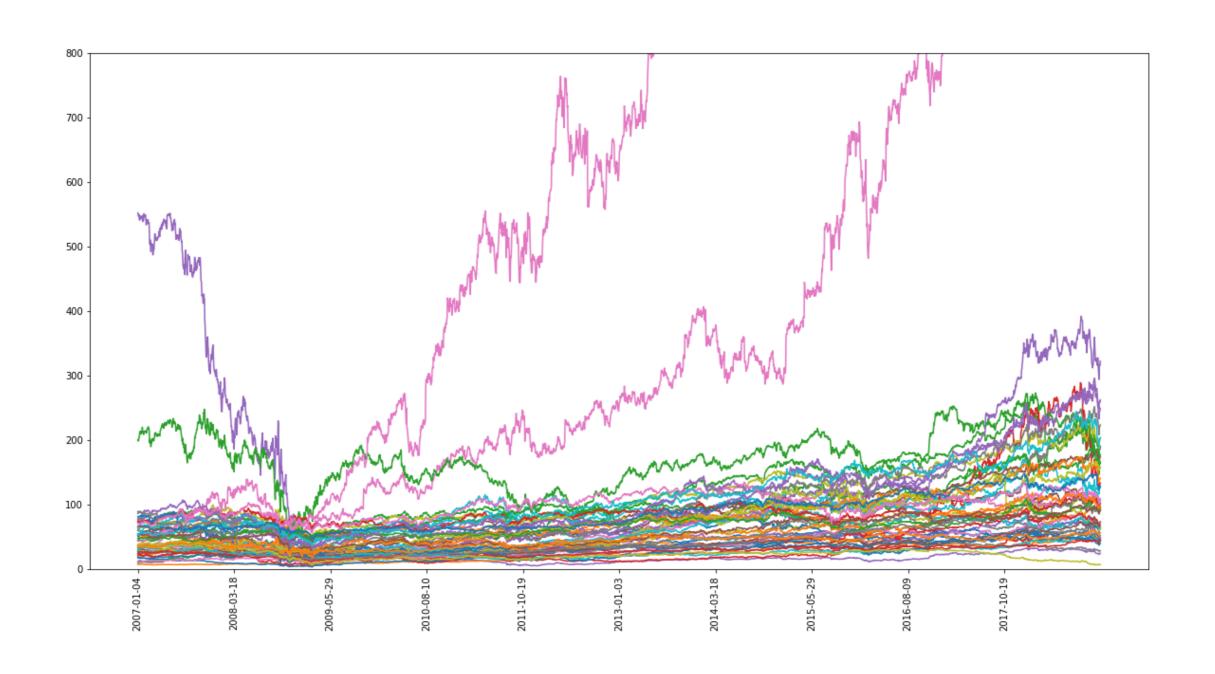


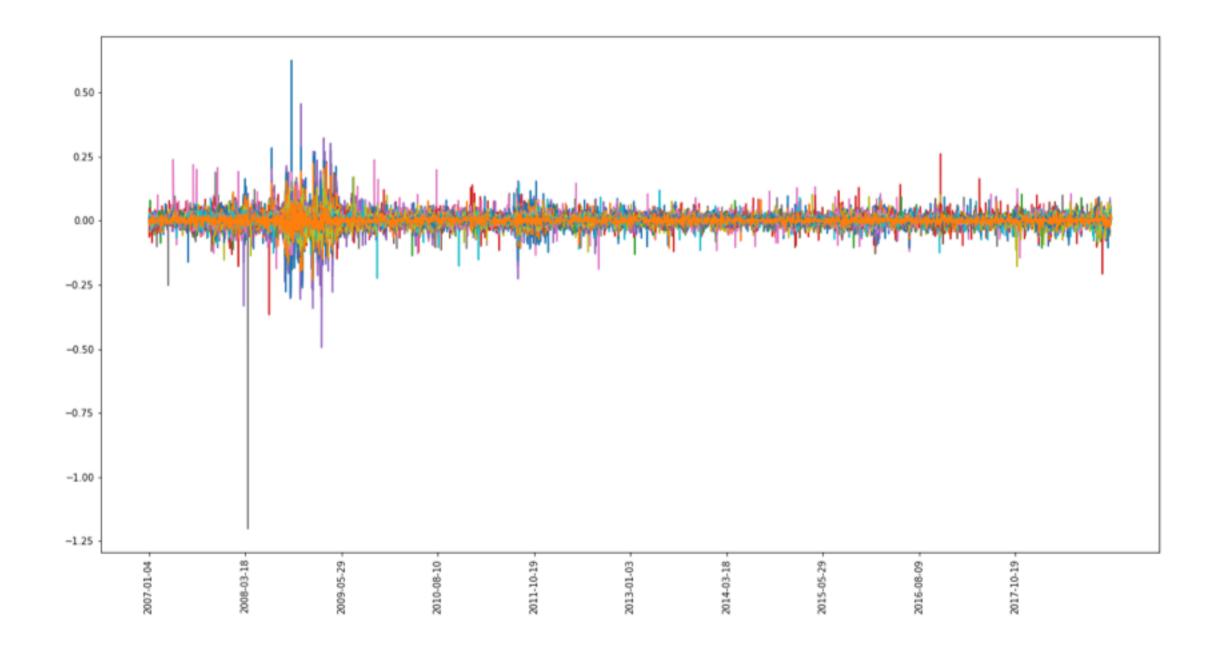


$$(1 - \epsilon)N(0_p, I_p) + \epsilon Q$$
$$N(5 * 1_p, I_p)$$

- Discriminator helps identify outliers or contaminated samples
- Generator fits uncontaminated portion of true samples

### Application: Price of 50 stocks from 2007/01 to 2018/12 Corps are selected by ranking in market capitalization





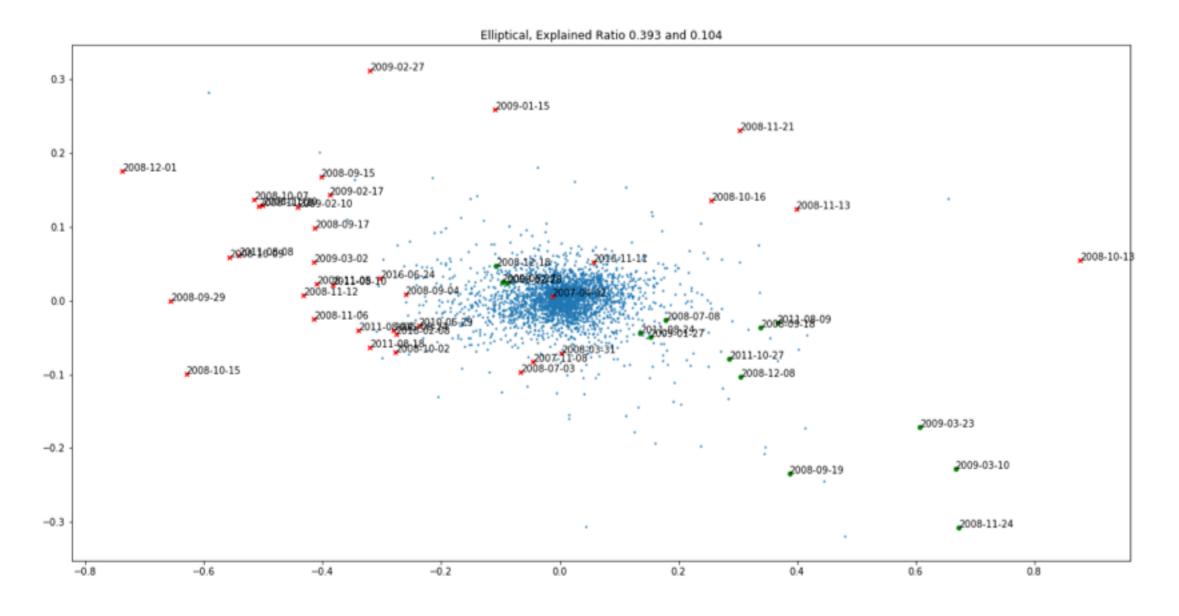
Log-return. y[i] = log(price\_{i+1}/price\_{i})

Fit data by Elliptica-GAN.

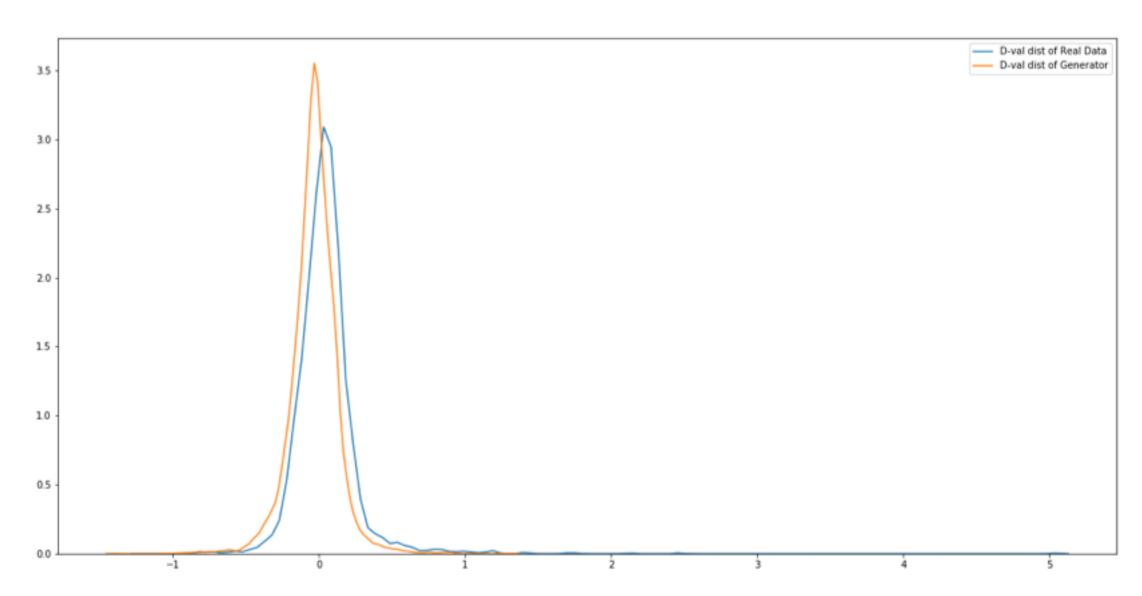
Apply SVD on scatter.

Dimension reduction on R^2.

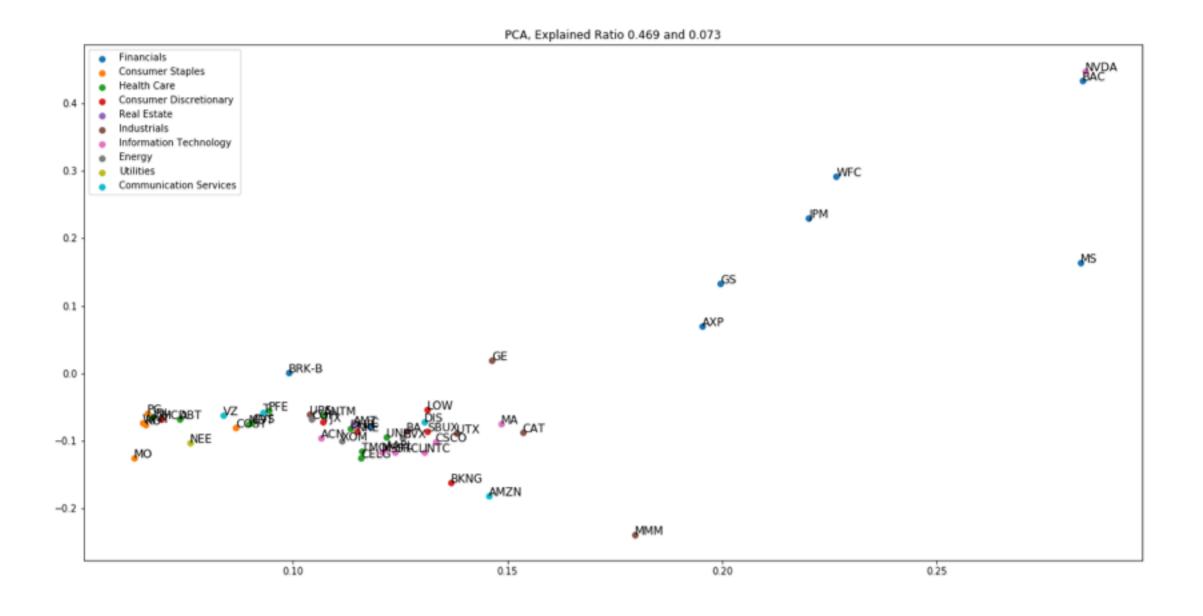




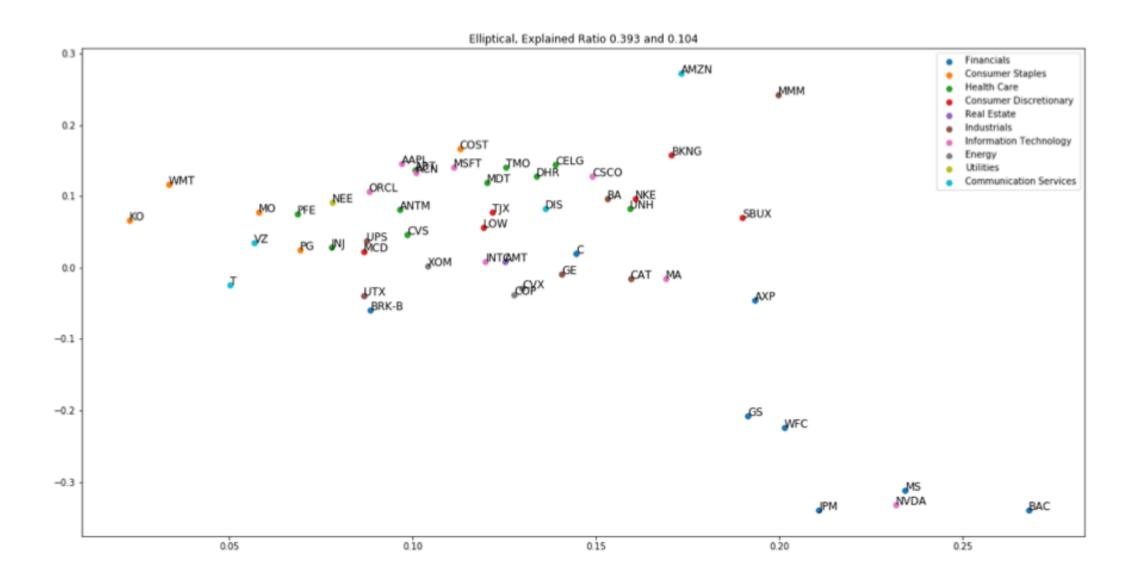
Discriminator value distribution from (Elliptical) Generator and real samples. Outliers are chosen from samples larger/ lower than a chosen percentile of Generator distribution



### Loading of PCA. First two direction are dominated by few corps —> not robust



#### Loading of Elliptical Scatter: Comparing with PCA, it's more robust in the sense that it does not totally dominate by Financial company (JPM, GS)



#### Reference

- Gao, Liu, Yao, Zhu, Robust Estimation and Generative Adversarial Networks, *ICLR 2019*, <a href="https://arxiv.org/abs/1810.02030">https://arxiv.org/abs/1810.02030</a>
- Gao, Yao, Zhu, Generative Adversarial Networks for Robust Scatter Estimation: A Proper Scoring Rule Perspective, <a href="https://arxiv.org/abs/1903.01944">https://arxiv.org/abs/1903.01944</a>

### Thank You

